

DATA MODELING USING TRANSMUTED DISTRIBUTIONS

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Abstract :

We discussed two distributions (the Lindley distribution and the converted Lindley distribution) and we dealt with some functions such as the pdf, cdf function, and we also discussed some statistical properties such as moments, the moment generation function, the coefficient of variation, flattening, torsion, and ordered statistics. IMSE The research concluded the importance of the converted Lindley distribution on the Lindley distribution and on all sample sizes in the practical aspect experimental .

Keywords: Lindley distribution ; Transmuted Lindley distribution; Probability Density Function ; Moments; Maximum likelihood estimation

1- Introduction:

Recently, several families of transformed distributions have been proposed, which are extended models with one or more parameters to the basic model of continuous distributions, both (Shaw and Buckley) used the rank transformation method (RTM), which is looking for a tool to build new families of non-distributives Gaussian, as it was used to modify a specific fundamental distribution for the purposes of modulating moments, especially skewness, torsion and kurtosis, introduced the QRTM quadratic rank transformation map, which many authors have used to introduce new important distributions such as the Lindley distribution transformed by (Nerovci) (2007) and the Lomax distribution Transformed by Lomax (1954), Where these new families provide more flexibility in modeling and analyzing real-life data in many applied fields, and this is why the statistical literature contains a good number of new families, where the distributions could not model some phenomena, and that is why they found shifting of the distributions in general or extending these distributions so that they may be compatible with some medical or biological phenomenon, In this article we have taken the Lindley distribution and used the transform map approach to define a new model named the Transformed Lindley Distribution (TLind), and according to the quadratic Rank transform map (QRTM). Approach the cumulative distribution function (CDF) satisfy the relationship

$$F_1(x)=(1+\lambda)F_2(x)-\lambda(F_2(x))^2$$

(1)

which on differentiation yields,

$$f_1(x) = f_2(x)[(1+\lambda) - 2\lambda F_2(x)] \quad (2)$$

where $F_2(x)$ is the CDF of the base distribution

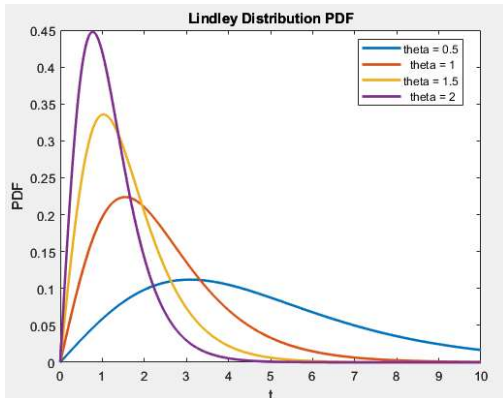
2-Transmuted Lindley Distribution (TLind)

The Lindley distribution describes the lifetimes of the processes and devices across a range of fields such as biology, engineering, and medicine. In fact, it's proven to be particularly effective in modeling mortality studies [3].

The (PDF) of Lindley Distribution is given by [2]:

$$f(x; \gamma) = \frac{\gamma^2}{\gamma + 1} (1 + x)e^{-\gamma x} \quad , \gamma > 0, x > 0 \tag{3}$$

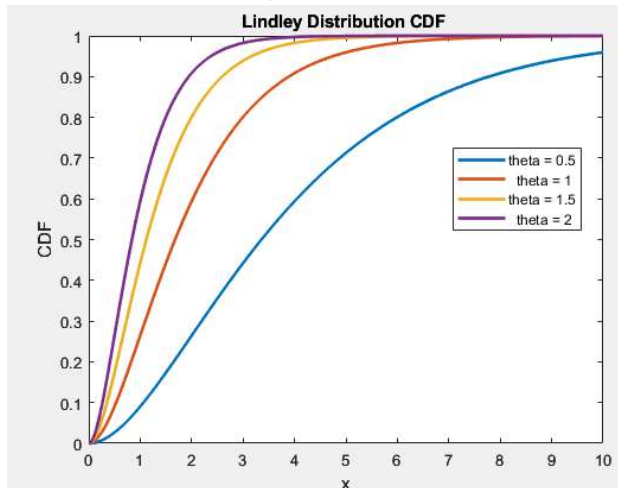
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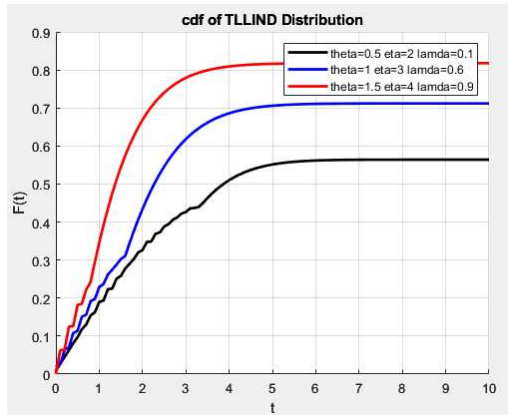
$$F(x; \gamma) = 1 - \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x} \quad , \gamma > 0, x > 0 \tag{4}$$



From equation (1) and (4) we conclude the (CDF) of the TLind [4].

$$F(x) = (1 - \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x}) (1 + \lambda \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x}) \tag{5}$$

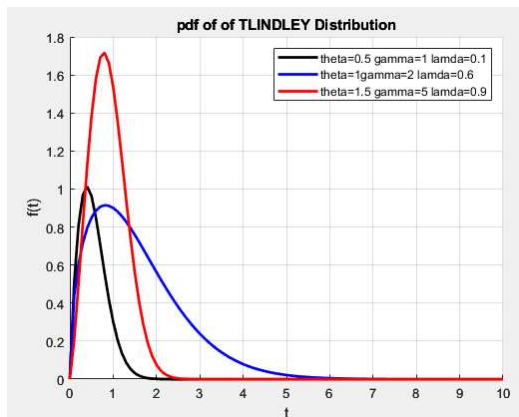
Where $|\lambda| \leq 1, \gamma > 0, x > 0$



The (PDF) of (TLind) random variable may be written as [4] :

$$f(x) = \frac{\gamma^2}{1 + \gamma} (1 + x) e^{-\gamma} \left(1 - \lambda + 2\lambda \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x} \right) \quad (6)$$

Where $|\lambda| \leq 1$, $\gamma > 0, x > 0$



3-Mathematical Properties of (TLind)

3.1 The Moments

Let x be a random variable that distributes the transmuted Lindley distribution that has the moment with parameter λ and $\theta^{[7]}$:

$$\begin{aligned} \mu_r &= E(X^r) = \int_{-\infty}^{\infty} x^r f(x) dx \\ E(X^r) &= \int_0^{\infty} x^r \frac{\gamma^2}{1 + \gamma} (1 + x) e^{-\gamma x} \left(1 - \lambda + 2\lambda \frac{(1 + \gamma + \gamma x)}{1 + \gamma} e^{-\gamma x} \right) dx \\ E(X^r) &= \frac{r!}{\gamma^r (\gamma + 1)} \left[(1 - \lambda)(\gamma + r + 1) + \frac{\lambda \gamma}{2^{r-1} (\gamma + 1)} (2\gamma + 3 + r) \right] \end{aligned} \quad (7)$$

The first moment of the TLind , when $r=1$, we get the following relationship :

$$\begin{aligned} E(X) = \mu &= \frac{1}{\gamma(\gamma+1)} \left[(1 - \lambda)(\gamma + 2) + \frac{\lambda \gamma}{(\gamma+1)} (2\gamma + 4) \right] \\ E(X) &= \frac{1}{\gamma(\gamma + 1)} \left[(1 - \lambda)(\gamma + 2) + \frac{2\lambda \gamma}{(\gamma + 1)} (\gamma + 2) \right] \end{aligned} \quad (8)$$

when $r = 2$ we get the second moment of the TLind .

$$E(X^2) = \frac{2}{\gamma^2(\gamma+1)} [(1-\lambda)(\gamma+3) + \frac{\gamma\lambda(2\gamma+5)}{2(\gamma+1)}] \tag{9}$$

The variance of the transmuted Lindley distribution is given by :

$$\text{Var}(x) = E(X^2) - (E(X))^2 = \frac{2}{\gamma^2(\gamma+1)} \left[(1-\lambda)(\gamma+3) + \frac{\gamma\lambda(2\gamma+5)}{2(\gamma+1)} \right] - \left[\frac{1}{\gamma(\gamma+1)} ((1-\lambda)(\gamma+2) + \frac{2\lambda}{\gamma+1}(\gamma+2)) \right]^2 \tag{10}$$

The Third moment of the TLind can be obtained by taking r = 3

$$E(X^3) = \frac{6}{\gamma^3(\gamma+1)} [(1-\lambda)(\gamma+4) + \frac{\gamma\lambda(\gamma+3)}{2(\gamma+1)}] \tag{11}$$

The Coefficient of Variation (CV) of transmuted Lindley distribution can be obtained by the following relationship :

$$CV = \frac{\sqrt{\text{Var}(x)}}{E(X)} = \frac{\sqrt{\frac{2}{\gamma^2(\gamma+1)} [(1-\lambda)(\gamma+3) + \frac{\gamma\lambda(2\gamma+5)}{2(\gamma+1)}] - \left[\frac{1}{\gamma(\gamma+1)} ((1-\lambda)(\gamma+2) + \frac{2\gamma\lambda(\gamma+2)}{(\gamma+1)}) \right]^2}}{\frac{1}{\gamma(\gamma+1)} [(1-\lambda)(\gamma+2) + \frac{2\gamma\lambda}{(\gamma+1)}(\gamma+2)]} \tag{12}$$

when r = 4 , the fourth moment of transmuted Lindley distribution is given :

$$E(X^4) = \frac{24}{\gamma^4(\gamma+1)} \left[(1-\lambda)(\gamma+5) + \frac{\gamma\lambda(2\gamma+7)}{8(\gamma+1)} \right] \tag{13}$$

As for the coefficient of Skewness (CS) and Kurtosis (CK) of the transmuted Lindley distribution can be obtained from :

$$\begin{aligned} CS &= \frac{E(X-\mu)^3}{(\text{Var}(x))^{3/2}} \\ E[(X-\mu)^3] &= E\left[\sum_{i=0}^3 \binom{3}{i} (-\mu)^i X^{3-i}\right] \\ &= E\left[\binom{3}{0} (-\mu)^0 X^3 + \binom{3}{1} (-\mu) X^2 + \binom{3}{2} (-\mu)^2 X + \binom{3}{3} (-\mu)^3 X^0\right] \\ &= E[X^3 - 3\mu X^2 + 3\mu^2 X - \mu^3] \\ &= E[X^3] - 3\mu E[X^2] + 3\mu^2 E[X] - \mu^3 \\ &= \mu_3 - 3\mu \mu_2 + 3\mu^2 \mu - \mu^3 \\ &= \mu_3 - 3\mu \mu_2 + 2\mu^3 \\ Sk &= \frac{\mu_3 - 3\mu \mu_2 + 2\mu^3}{(\text{Var}(x))^{3/2}} \end{aligned} \tag{14}$$

As for kurtosis of TLind , it is written as follows :

$$CK = \frac{E(X-\mu)^4}{(\text{Var}(x))^2}$$

$$\begin{aligned}
 E[(X - \mu)^4] &= E\left[\sum_{i=0}^4 \binom{4}{i} (-\mu)^i X^{4-i}\right] \\
 &= E\left[\binom{4}{0}(-\mu)^0 X^4 + \binom{4}{1}(-\mu)X^3 + \binom{4}{2}(-\mu)^2 X^2 + \binom{4}{3}(-\mu)^3 X^1 + \binom{4}{4}(-\mu)^4 X^0\right] \\
 &= E[X^4 - 4\mu X^3 + 6\mu^2 X^2 - 4\mu^3 X + \mu^4] \\
 &= E[X^4] - 4\mu E[X^3] + 6\mu^2 E[X^2] - 4\mu^3 E[X] + \mu^4 \\
 &= \mu_4 - 4\mu \mu_3 + 6\mu^2 \mu_2 - 4\mu^3 \mu + \mu^4 \\
 &= \mu_4 - 4\mu \mu_3 + 6\mu^2 \mu_2 - 3\mu^4 \\
 CK &= \frac{\mu_4 - 4\mu \mu_3 + 6\mu^2 \mu_2 - 3\mu^4}{(Var(x))^2} \tag{15}
 \end{aligned}$$

3.2 Moment Generating Function

The moment generation function (MGF) of the TLind is ^[6] :

$$M_X(t) = E[e^{xt}]$$

$$\begin{aligned}
 M_X(t) &= \int_0^\infty e^{tx} f(x) dx \\
 &\text{where } f(x) \text{ is (PDF) of transmuted Lindley distribution.} \\
 M_X(t) &= \int_0^\infty e^{tx} \frac{\gamma^2}{1+\gamma} (1+x) e^{-\gamma x} (1-\lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma}) e^{-\gamma x} dx \\
 M_X(t) &= \frac{\gamma^2}{1+\gamma} \int_0^\infty e^{(t-\gamma)x} (1+x) (1-\lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma}) e^{-\gamma x} dx \tag{16}
 \end{aligned}$$

Where $x > 0, \gamma > 0, |\lambda| \leq 1$

3.3 Survival Function

The survival function $S(x)$ for the TLind is given by :

$$\begin{aligned}
 S(x) &= 1 - F(x) \\
 S(x) &= 1 - \left[\left(1 - \frac{(1+\gamma+\gamma)}{1+\gamma} e^{-\gamma x}\right) \left(1 + \lambda \frac{(1+\gamma+\gamma)}{1+\gamma} e^{-\gamma x}\right) \right] \\
 S(x) &= \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \left(1 - \lambda + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma} \right) \quad x > 0, \gamma > 0, \lambda > 0 \tag{17}
 \end{aligned}$$

3.4 Order Statistics

Let X_1, X_2, \dots, X_n are a simple random variables with distribution function (CDF) $F_X(x)$ and (PDF) $f_X(x)$. The corresponding order statistics are the X_i arranged in ascending order. The largest of the X_i is denoted by $X_{(n)}$ and the smallest of the X_i is denoted by $X_{(1)}$ ^[4], the PDF of $X_{(i)}$ is given by ^[11]:

$$f_{X(i)}(x) = \frac{n!}{(i-1)!(n-i)!} f_X(x) [F_X(x)]^{i-1} [1 - F_X(x)]^{n-i} \tag{18}$$

The PDF of order statistics for the TLind is given by ^[7] :

$$\begin{aligned}
 & f_{X(i)}(x) \\
 &= \frac{n!}{(i-1)!(n-i)!} \left(\frac{\gamma^2}{\gamma+1} (1+x)e^{-\gamma x} (1-\lambda + 2\lambda \frac{1+\gamma+\gamma x}{\gamma+1} e^{-\gamma x}) \right) \left[\left(1 - \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \left(1 + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \right]^{i-1} \left[\frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} (1-\lambda + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x}) \right]^{n-i} \tag{19}
 \end{aligned}$$

As for The PDF of the largest order statistic $X_{(n)}$ is given by^[7]:

$$f_{X(n)}(x) = \frac{n\gamma^2}{\gamma+1} (1+x)e^{-\gamma x} (1-\lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x}) \left[\left(1 - \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \left(1 + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \right]^{n-1} \tag{20}$$

And the PDF of the smallest order statistic $X_{(1)}$ is given by^[7]:

$$\begin{aligned}
 & f_{X(1)}(x) \\
 &= \frac{n\gamma^2}{1+\gamma} (1+x)e^{-\gamma x} (1-\lambda + 2\lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x}) \left[1 - \left(1 - \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \left(1 + \lambda \frac{1+\gamma+\gamma x}{1+\gamma} e^{-\gamma x} \right) \right]^{n-1} \tag{21}
 \end{aligned}$$

4-Estimation Method of (TLind)

We used the Maximum Likelihood Method for estimating the parameters of TLind . Let X_1, X_2, \dots, X_n is the sample size n at random from the TLind .Then the likelihood function You write as follows ^[7]:

$$L = (x_i, \gamma, \lambda) = \prod_{i=1}^n f(x_i, \gamma, \lambda)$$

$$L = \prod_{i=1}^n \left[\frac{\gamma^2}{\gamma+1} (1+x)e^{-\gamma x} \left(1 - \lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma} \right) \right]$$

$$L = \frac{\gamma^{2n}}{(\gamma+1)^n} e^{-\sum_{i=1}^n \gamma x_i} \prod_{i=1}^n (1+x_i) \cdot (1-\lambda + 2\lambda \frac{(1+\gamma+\gamma x)}{1+\gamma} e^{-\gamma x})$$

$$\ln L = 2n \ln \gamma - n \ln(\gamma+1) - \sum_{i=1}^n \gamma x_i + \sum_{i=1}^n \ln(1+x_i) + \sum_{i=1}^n \ln \left(1 - \lambda + 2\lambda \frac{1+\gamma+\gamma x_i}{\gamma+1} e^{-\gamma x_i} \right)$$

$$\frac{\partial L}{\partial \gamma} = \frac{2n}{\gamma} - \frac{n}{\gamma+1} - \sum_{i=1}^n x_i + \sum_{i=1}^n \frac{2\lambda x_i \frac{1+\gamma+\gamma x_i}{1+\gamma} e^{-\gamma x_i} e^{-\gamma x_i} \left[\frac{1}{(\gamma+1)^2} - \frac{1+\gamma+\gamma x_i}{1+\gamma} \right]}{1-\lambda + 2\lambda \frac{1+\gamma+\gamma x_i}{1+\gamma} e^{-\gamma x_i}} = 0$$

$$\frac{\partial L}{\partial \lambda} = \sum_{i=1}^n \frac{2^{\frac{1+\gamma+\gamma x_i}{\gamma+1}} e^{-\gamma x_i} - 1}{1 - \lambda + 2\lambda^{\frac{1+\gamma+\gamma x_i}{\gamma+1}} e^{-\gamma x_i}} = 0$$

And using quasi-Newtonian algorithm to maximize the incoming registration probability functio

5-Simulation Concept

We conducted Monte Carlo simulation studies to assess on the finite sample behavior of the maximum likelihood and the simulations were carried out using the statistical software Matlab . 2022b program The true parameter values used in the data generating processes are $\gamma = 2, \lambda = 0.1$ and sample sizes $n=25, 50, 75, 100, 150$.

Assuming that the default values for the parameters of Lind is $\gamma = 2$ and TLind is $\gamma = 2, \lambda = 0.1$, the estimation results can be summarized as shown in Table (1).

Table (1): The values of the real density function and estimated according to the method of MLE of the distribution of Lindley, and TLindley at the theoretical samples sizes :

N	$f(x)_{Li}$	$\hat{f}(x)_{Li}$	MSE	$f(x)_{TL}$	$\hat{f}(x)_{TL}$	MSE	
25	0.924	0.659	0.070	0.898	0.860	0.00146	
	74	81	19	51	29		
	0.734	0.589	0.021	0.881	0.845		0.00130
	60	34	10	73	63		
	0.729	0.587	0.020	0.845	0.813		0.00102
	27	17	19	43	57		
	0.702	0.576	0.016	0.719	0.700		0.00038
	75	18	02	74	12		
	0.658	0.557	0.010	0.696	0.678		0.00031
	65	23	29	49	83		
0.566	0.514	0.008	0.695	0.678	0.00031		
62	78	39	80	20			
0.494	0.478	0.008	0.624	0.612	0.00015		
06	17	13	53	45			
0.423	0.439	0.006	0.592	0.582	0.00010		
76	56	71	24	44			
0.417	0.436	0.002	0.532	0.526	0.00004		
64	04	69	33	39			
0.172	0.262	0.001	0.441	0.440	0.00004		
55	71	72	51	57			
IMSE			0.239			0.00021	
			36				
AIC			-			-54.27134	
			12.38				

			310			
Best			$\hat{f}(x)_{TLind}$			
50	0.859 14	0.572 41	0.082 21	0.625 46	0.585 35	0.00161
	0.821 67	0.562 85	0.066 99	0.549 86	0.525 51	0.00059
	0.813 47	0.560 69	0.063 90	0.449 10	0.442 60	0.00044
	0.613 30	0.500 01	0.012 83	0.446 40	0.440 33	0.00044
	0.610 05	0.498 89	0.012 36	0.441 90	0.436 54	0.00044
	0.601 79	0.496 00	0.011 96	0.441 88	0.436 52	0.00043
	0.546 79	0.475 93	0.011 94	0.433 98	0.429 84	0.00043
	0.476 36	0.447 79	0.011 81	0.413 45	0.412 38	0.00042
	0.462 08	0.441 71	0.011 74	0.410 73	0.410 06	0.00042
	0.445 92	0.434 66	0.011 46	0.398 84	0.399 86	0.00042
IMSE			0.122 95			0.00021
AIC			- 26.75 951			-58.56444
Best			$\hat{f}(x)_{TLind}$			
75	0.731 24	0.560 55	0.029 13	0.644 37	0.675 75	0.00098
	0.700 32	0.549 75	0.022 67	0.638 48	0.668 71	0.00091
	0.685 59	0.544 46	0.019 92	0.631 65	0.660 56	0.00084
	0.674 08	0.540 26	0.017 91	0.627 00	0.655 02	0.00079
	0.638 65	0.526 97	0.012 47	0.606 61	0.630 84	0.00059
	0.626 23	0.522 18	0.010 83	0.592 47	0.614 15	0.00047
	0.609 47	0.515 59	0.010 44	0.563 76	0.580 55	0.00029

	0.607 92	0.514 98	0.010 42	0.549 47	0.563 93	0.00029
	0.595 29	0.509 91	0.010 04	0.518 68	0.528 40	0.00029
	0.547 11	0.489 85	0.009 88	0.515 83	0.525 13	0.00028
IMSE			0.094 71			0.00017
AIC			- 41.15 213			-90.78444
Best	$\hat{f}(x)_{TLind}$					
100	0.704 75	0.577 77	0.016 12	0.535 66	0.554 02	0.00086
	0.695 50	0.573 86	0.014 80	0.531 10	0.548 27	0.00086
	0.685 12	0.569 42	0.013 39	0.508 68	0.520 18	0.00085
	0.624 37	0.542 49	0.008 34	0.496 33	0.504 86	0.00084
	0.622 63	0.541 70	0.008 18	0.477 64	0.481 88	0.00084
	0.612 22	0.536 90	0.007 81	0.471 03	0.473 80	0.00083
	0.583 19	0.523 25	0.007 76	0.460 26	0.460 71	0.00083
	0.565 16	0.514 54	0.007 07	0.459 40	0.459 68	0.00082
	0.556 74	0.510 42	0.006 70	0.457 99	0.457 97	0.00078
	0.554 03	0.509 09	0.006 55	0.444 95	0.442 26	0.00076
IMSE			0.079 09			0.00017
AIC			- 55.51 464			-110.78886
Best	$\hat{f}(x)_{TLind}$					
150	0.573 78	0.525 86	0.007 46	0.477 42	0.477 60	0.00000
	0.572 99	0.525 45	0.007 45	0.475 56	0.475 74	0.00000

	0.562 40	0.519 98	0.007 44	0.475 47	0.475 65	0.00000
	0.556 17	0.516 72	0.007 43	0.466 93	0.467 07	0.00000
	0.532 35	0.504 11	0.007 38	0.457 69	0.457 80	0.00000
	0.525 97	0.500 67	0.007 28	0.451 97	0.452 06	0.00000
	0.489 59	0.480 64	0.007 26	0.447 01	0.447 08	0.00000
	0.480 49	0.475 51	0.007 22	0.445 77	0.445 83	0.00000
	0.469 73	0.469 37	0.007 22	0.425 70	0.425 68	0.00000
	0.573 78	0.525 86	0.007 46	0.416 00	0.415 95	0.00000
IMSE			0.060 23			0.00001
AIC			- 84.30 144			-212.54533
Best	$\hat{f}(x)_{TLind}$					

Table(1)showed that under theoretical values of parameters of distributions ($\theta=2$ of Lindley distribution and $(\theta=2, \lambda=0.1)$ of transmuted Lindley distribution the following: under $n=25, 50, 75, 100, 150$ the transmuted Lindley distribution is the best of other distributions because the IMSE and AIC for $\hat{f}(x)_{TLind}$ less than the Lindley distribution for all sizes

6- Conclusion

The efficiency of the transmuted distributions proved, as we noticed through the results that the transmuted Lindley distribution reconciled with the original distribution

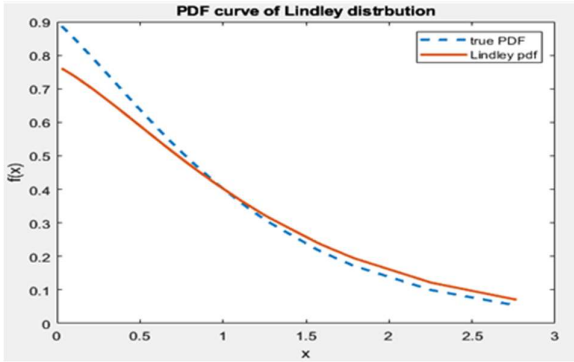


Figure (1.1) Real and Maximum Likelihood curve for Lindley distribution under $n=25$

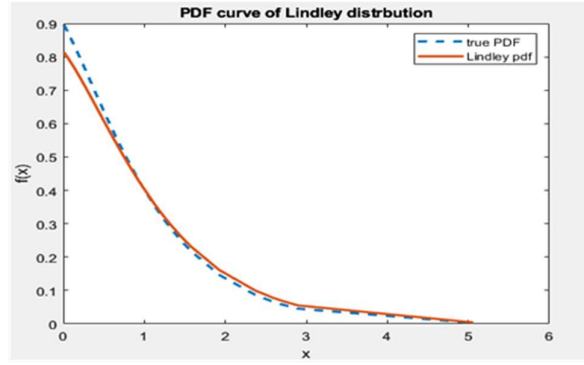


Figure (1.2) Real and Maximum Likelihood curve for Lindley distribution under $n=50$

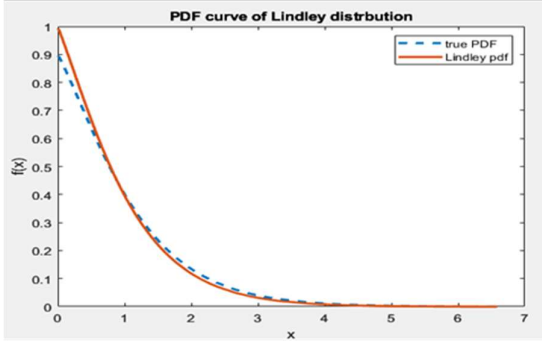


Figure (1.3) Real and Maximum Likelihood curve for Lindley distribution under $n=100$

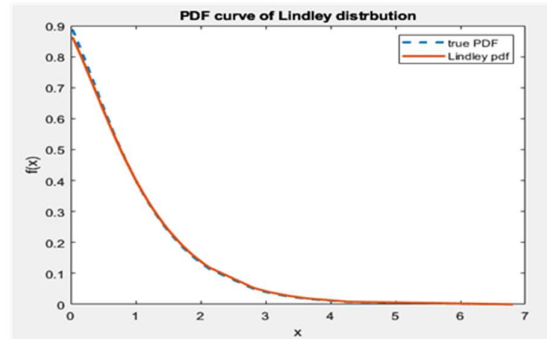


Figure (1.4) Real and Maximum Likelihood curve for Lindley distribution under $n=75$

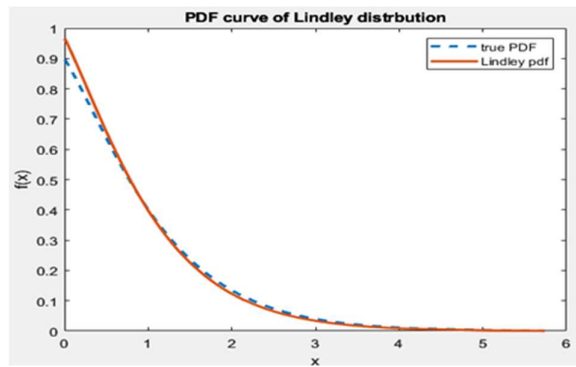


Figure (1.5) Real and Maximum Likelihood curve for Lindley distribution under $n=150$

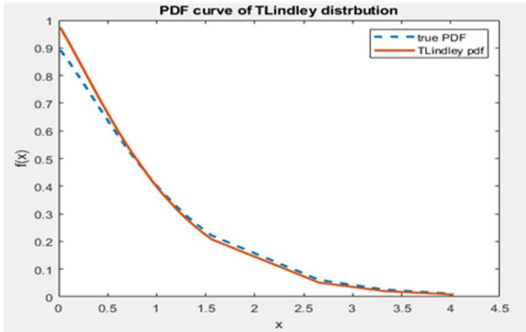


Figure (1.6) Real and Maximum Likelihood curve for TLindley distribution under n=25

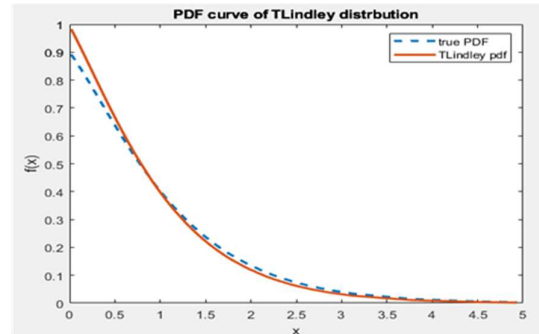


Figure (1.7) Real and Maximum Likelihood curve for TLindley distribution under n=50

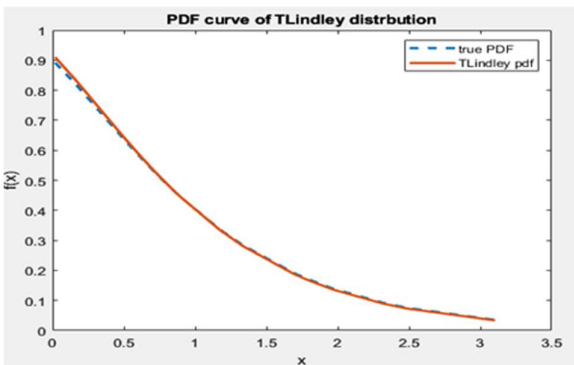


Figure (1.8) Real and Maximum Likelihood curve for TLindley distribution under n=75

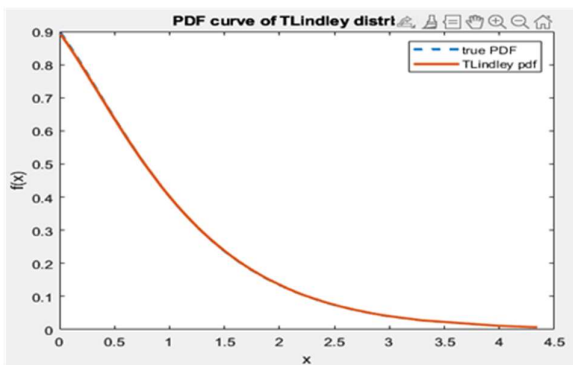


Figure (1.9) Real and Maximum Likelihood curve for TLindley distribution under n=100

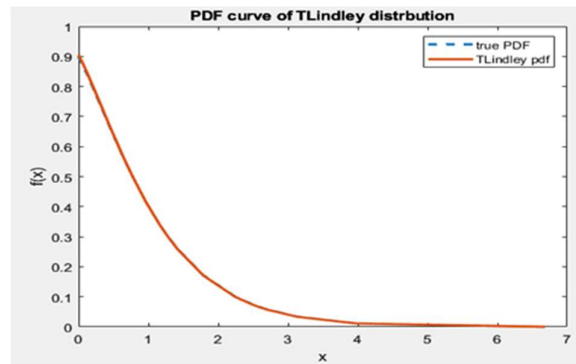


Figure (1.10) Real and Maximum Likelihood curve for TLindley distribution under n=150

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